



The local innovation spillovers of listed firms[☆]

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ABSTRACT

This paper provides evidence of local innovation spillovers (i.e., innovation by one firm fostering innovation by neighboring firms). First, I document that exogenous shocks to innovation by listed firms affect innovation by private firms in the same geographical area and that such local innovation spillovers decline rapidly with distance. Second, these local innovation spillovers stem from knowledge diffusing locally through two channels: learning across local firms and inventors moving from their employer to both existing firms and newly started spin-outs. Finally, I study the two-way relations between innovation spillovers and the availability of capital. I find that local innovation spillovers cause venture capital funds from outside the area to invest more in the local area, and that capital availability amplifies local innovation spillovers.

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1. Introduction

The success of innovation clusters such as Silicon Valley is often explained by local networks of innovative firms helping to diffuse knowledge across firms. It has motivated large investments by governments to promote such clusters. Often, particular emphasis is put on developing ecosystems of large and small firms, such as the recent American “Regional Cluster Initiative” funded by the Economic Development and Small Business administrations. Supporters of such policies stress that the knowledge

produced by large firms will benefit neighboring smaller firms, as exemplified by Seattle’s innovation cluster that began developing after Microsoft relocated its headquarters to the area.

The spatial concentration of innovative activities is expected to foster innovation because, as for economic spillovers in general, agglomeration allows local firms to share inputs, workers, and ideas more efficiently.¹ However, while strong evidence exists that agglomeration and innovation are correlated, causal identification remains elusive as innovation trends for all firms located in the same area are likely driven by the same underlying local factors, such as leading research universities, benign weather conditions, and tax advantages (Carlino and Kerr, 2015).

To disentangle innovation spillovers from the effects of local conditions, I exploit a shock on the research labs of listed firms in a given geographical area produced by a reg-

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¹ Surveys about the link between knowledge and agglomeration include Moretti (2004) or Carlino and Kerr (2015).

ulatory change coming from a different state. The shock is caused by the staggered adoption by individual states of business combination (BC) laws, preventing acquirers from using the target's assets to pay down acquisition debt. The laws make it more difficult to complete hostile takeovers of listed firms incorporated in the adopting state. The lower takeover threat has been shown to have weakened external governance, allowing management to enjoy “the quiet life” (Bertrand and Mullainathan, 2003), resulting in a decrease in innovation by listed firms, even in areas outside their state of incorporation.

BC laws provide an appealing shock to identify innovation spillovers because they cause local areas to experience variation in the activities of listed firms' research labs driven by out-of-state shocks, while not having a direct effect on the innovation of local private firms. My hypothesis is that changes in listed firms' innovation in a given area directly affects innovation by private firms in the same area. I limit the concerns that local private firms may be affected by other changes in the state correlated with the adoption of business combination laws by focusing on the innovation activity of listed firms outside their state of incorporation and by controlling for important economic characteristics at the local area level.

I study innovation by US firms over the 1975–2000 period. I use patent and inventor data from the USPTO containing information about patent inventors, including addresses and employers. Inventor addresses allow me to allocate innovations to different commuting zones (i.e., local geographic areas encompassing all metropolitan and non-metropolitan areas in the US. I consider that a firm is active in a commuting zone if it files patents in that area). The dataset covers both listed and private firms, a classification my identification strategy exploits. In the data, both sets of firms account for a similar fraction of patents filed: around 60% for listed firms and 40% for private firms. However, as one would expect, they differ in the degree of the geographical concentration of their patenting activity. On average, listed firms produce patents in 12 different commuting zones relative to 1.5 for private firms. Moreover, listed firms produce less than 20% of their patents in their state of incorporation.

The activity of listed firms' research labs in a specific area will be driven by out-of-state shocks at different points in time, which depends on when a given state adopts a BC law. Moreover, the same shock affects different areas with different intensities, depending on how many listed firms active in an area are incorporated in the adopting state. This allows me to employ a difference-in-differences strategy to study how innovation by private firms reacts to the change in innovation of listed firms' research labs in the same area, triggered by the adoption of out-of-state BC laws.

In the first part of the paper, I study both how large these spillovers are and how local they are. I find that the out-of-state shocks on listed firms generate negative and economically significant spillovers onto private firms in the same area and that this result is robust to the inclusion of controls for commuting zone-level innovation capacities and labor characteristics as well as to sample restrictions,

such as excluding the most innovative cities or states. I also find these innovation spillovers to be markedly local (i.e., they fade away quickly with distance). Indeed, shocks on listed firms' research labs in a given commuting zone have spillovers mostly for private firms in the same commuting zone. For private firms' innovation in other commuting zones within 100 miles, spillovers are still negative, but small: the elasticity is divided by a factor of three. Beyond 100 miles, spillovers are indistinguishable from zero.

To obtain an economic interpretation of the reduced form shock on listed firms, I predict the number of patents produced by listed firms in the area using the BC shock and estimate an elasticity of 0.2, which implies that decreasing innovation by listed firms in a given area will reduce innovation by private firms in the same area by more than 20%.

Next, I test whether the spillovers of listed firms onto private firms can be explained by an overall change in listed firm efficiency or if private firms change their innovation policy in reaction to a change in listed firms' innovation. I find limited support for the overall efficiency hypothesis. In particular, I find no local spillovers onto non-innovative firms and no change in the probability for private innovative firms being acquired. I also find that private firms' innovation displays the same changes whether firms are in the upstream or downstream industries and whether they are or are not suppliers of listed firms. This makes it unlikely that the identified spillovers are driven by a demand channel, whereby in response to BC laws, listed firms would change their demand for technologies that would affect innovation by private firms.

The changes in innovation by private firms appears to be driven by a variation in innovation by listed firms, rather than a change in other listed firm policies. I explore the two channels specific to these “*innovation spillovers*”: a direct competition in the market for ideas and the existence of knowledge spillovers between listed and private firms when they are in the same area. I find limited evidence that changes in private firms' innovation are driven by a change in their incentives to innovate in reaction to variation in local competition with listed firms. By contrast, I find that private firms even in very unrelated industries (e.g., computer vs. chemical) still strongly benefit from the knowledge created by listed firms, consistent with the fact that innovation can stimulate the production of other new ideas across sectors.

Since knowledge spillovers appear to be the main driver behind the existence of the broader innovation spillovers, I study in detail which pipes knowledge can go through to spread across local firms and focus on two main channels: learning from local firms and inventors moving from their employer to both existing and newly started local firms. First, I find evidence of knowledge diffusion via learning across local firms. Indeed, I document higher local innovation spillovers onto firms that are technologically closer to the listed firms innovating locally (those that file patents in the same technological classes, or tend to cite patents filed by listed firms or by local firms). For each proxy of technological proximity, I find that a one standard deviation of this proxy amplifies local innovation spillovers by about half of the average effect.

Knowledge diffusion through learning across local firms is also likely to depend on the local supply of educated workers, whose ability to incorporate and apply new knowledge may be more important than for low-skilled workers (e.g., [Moretti, 2004](#)). For each commuting zone, I calculate the supply of college graduate workers at the beginning of the period. I find that commuting zones in the 75th percentile of the distribution of educated workers experience local innovation spillovers that are twice as large as those for commuting zones in the 25th percentile. I find similar results using instruments that exploit historical differences in the supply of colleges to predict the fraction of educated workers today.

Second, I find evidence of knowledge diffusion via employees moving across local firms. In a first test, I exploit variation across states in the enforcement of non-compete clauses that limit worker mobility. I find that commuting zones in states that allow for more mobility experience local innovation spillovers that are twice as large as those in states that do not.

I also study how variation in regulation-induced changes in listed firms' innovation affects the mobility of inventors from listed firms to both existing private firms and newly started spin-outs in the same area. I define a spin-out as a new firm employing, in the first year it files patents, inventors formerly employed by a listed firm active in the same area. For both existing and new firms, I observe more mobility when the stock of patents by listed firms increases.

I investigate the two-way connection between local innovation spillovers and the availability of venture capital. First, I examine whether local innovation spillovers attract capital to the area. To identify non-local investors, I use the VentureXpert database, which reports for each venture capital (VC) fund covered its address and the location of all its investments. I find that when listed firms in a commuting zone innovate more, VC funds located outside that commuting zone increase the volume of their investments in that commuting zone. On average, a one standard deviation increase in the stock of patents by listed firms in a commuting zone increases non-local VC investments per year by 11%. This is all the more remarkable given that non-local investments are rare in the VC industry.

Second, I test whether conversely, exogenous fluctuations in local capital availability amplify local innovation spillovers by enabling local firms to better finance innovations. To do so, I instrument the amount of VC capital available locally using variation in the size of state pension funds. Because state pension funds invest disproportionately in local investment funds, such as private equity and venture capital funds, local investment funds raise capital more easily when local pension pools are larger ([Gonzalez-Urbe, 2020](#)). I find that commuting zones in the 75th percentile in the distribution of (exogenous) VC financing experience local innovation spillovers that are twice as large as those in commuting zones in the 25th percentile.

Taken together, the paper shows that sizeable local innovation spillovers exist, and are at least partly driven by knowledge diffusion via learning across local firms, as well as employees and inventors moving across local firms. Fur-

thermore, these spillovers attract capital to the area, which amplifies the spillovers. These findings point to several policy implications. If the clustering of innovation were mostly due to attractive local attributes (universities, etc.), local public policies aimed at fostering innovation clusters should focus on providing those. However, if instead, innovation clusters stem from innovation spillovers, then subsidies can be justified. My findings also suggest that local innovation spillovers can be amplified by policies promoting intrastate labor mobility, by restricting non-compete clauses, by improving the supply of skilled labor (e.g., via the construction of college institutions), and by improving access to capital.

Literature Review. This paper contributes to several strand of literature. First, it relates to studies examining how the stock of external knowledge available in the surroundings of economic agents affects their productivity and ability to innovate. The dominant approach in this literature is to regress productivity, wages (used as a proxy for productivity) or innovation on a proxy for the stock of knowledge available, such as the stock of R&D (e.g., [Peri, 2005](#)), the supply of college graduates (e.g., [Rauch, 1993](#)), population density (e.g., [Ciccone and Hall, 1996](#)) or firm density (e.g., [Greenstone et al., 2010](#); [Guiso and Schivardi, 2011](#)).

Second, my paper relates to the literature studying how corporate investment is shaped by the firm's neighbors. This question has been studied for investment in general (e.g., [Dessaint et al., 2018](#)), as well as for firm creation (e.g., [Guiso et al., 2021](#)) and innovations in particular ([Peri, 2005](#); [Bloom et al., 2013](#)). I add to this literature by providing a new method of studying innovation spillovers and by providing evidence for specific channels through which these local innovation spillovers can occur. I also use a finer measure of geographic proximity by using inventor addresses rather than firms' headquarters as the location of innovation, as it is not clear that information regarding failed or successful innovative projects will be communicated by CEOs. Inventors appear more likely to spread knowledge locally, in particular by moving across firms. Finally, I study the specific interactions between publicly listed and private firms, which is a subject that has received little attention thus far.

More broadly, my paper relates to the literature on urban economics literature and agglomeration. Most of these studies focus on fixed characteristics, while a few recent papers are exploring how finance can produce changes in these agglomeration forces both within countries [e.g., [Hombert and Matray, 2017](#) for the US, [Bau and Matray, 2020](#) for India], within sectors ([Hombert and Matray, 2020](#)), and across countries ([Xu, 2020](#)).

In this burgeoning literature, two very different mechanisms are at play that explain the comovement of behaviors. The first mechanism is that managers either infer information from their peers or simply "mimic" these peers, and the second is that neighboring firms have a direct effect on their peers' inputs or cash-flows. My paper is about the second mechanism. I show that innovation by private firms is affected because a key input in their own innovation production functions varies: the local stock of external knowledge produced by listed firms.

2. Data

2.1. Innovation

I use patents filed with the US Patent and Trademark Office (USPTO), as compiled in the National Bureau of Economic Research (NBER) Patents File (Hall et al., 2002) to measure innovation. These data contain all patents granted in the US, including information about the patentee (including a unique identifier, institutional characteristics, nationality, and geographic location) and the patent (year of application, technology class, and number of citations received). An appealing feature of the NBER Patents File is that it covers the entire universe of patents filed in the US, including patents filed by young and private firms.

Both listed firms and private firms play an important role in innovation activity in the US. Throughout my sample period, the fraction of patents filed by listed firms is relatively stable at approximately 50–60%.

I keep only those patents filed by US corporations in my sample and exclude patents filed by foreign firms, universities, and government agencies. I date patents by the year in which the application was filed to avoid anomalies resulting from a lag between the application and grant dates. I consider all patents filed between 1975 and 2000 (the first year and last year where the truncation bias is limited).

To obtain the location of the inventors at the county level, I use the Harvard Patent Database,² which provides the latitude and longitude for each inventor associated with a patent. These coordinates can then be used to obtain the exact county in which a patent was developed (Hombert and Matray, 2018).

2.2. Geographic area: commuting zones

Commuting zones are 741 clusters of counties that are characterized by strong commuting ties within commuting zones and weak commuting ties across commuting zones. I restrict my analysis to commuting zones in which I can observe at least one patent during the 1975–2000 period, which results in a balanced panel of 685 distinct commuting zones, mapping to 48 states in the US (missing are Alaska, Hawaii, and the District of Columbia).

Commuting zones have two main advantages. First, they are based on economic ties rather than political boundaries and, as such, are a more suitable candidate for estimating the scope of innovation spillovers. Indeed, they are sufficiently small so that spillovers can plausibly occur (as knowledge spillovers tend to occur on relatively small scales) and their geographical boundaries can be defined in a constant way over time, allowing the analysis over a long time period. Second, they cover the entire US (as opposed for instance to metropolitan statistical areas, which captures only a third of all counties in the US).

To measure the existence of geographical spillovers, I aggregate patents at the commuting zone level. This is motivating by the fact that innovation can trigger the pro-

duction of other new ideas across sectors. For instance, Jaffe et al. (1993) report that up to 25% of citations occur across five broad technological fields. When looking at the three digit level (approximately 450 technological fields) approximately 40% of citations are across fields. Aggregating at the commuting zone level allows me to capture these potential cross-sector spillovers.

2.3. Local labor market characteristics

I construct different characteristics at the commuting zone level using various data sources. The main source is the Census Integrated Public Use Micro Samples for the years 1970, 1980, 1990, and 2000 (Ruggles et al., 2010). I apply the usual restrictions to compute labor market characteristics: individuals must be between the ages of 16 and 64 and be working in the year preceding the survey and I drop residents of institutional groups such as prisons and psychiatric institutions, as well as unpaid family workers. Population estimates on a yearly basis are from the Census. Appendix A.1 details the construction of the variables.

Data on venture capital activity and the availability of venture capital funds come from the VentureXpert database. I identify the commuting zone in which the fund is located and where it makes an investment using the ZIP code information provided by Venture Xpert. Finally, data regarding educational attainment, number of colleges, and federal R&D expenses are from the National Science Foundation's CASPAR database. Table 1 provides summary statistics for the main variables.

3. Identification strategy

3.1. Empirical specification

To test the existence of innovation spillovers from listed firms onto private firms, one would ideally like to regress innovation by private firms on innovation by listed firms in the same area. The main challenge when doing so is that the innovation activity of both private and listed firms in a given city is likely to be determined by common specific location factors, such as the proximity with universities or the quality of local amenities. Therefore, it is also quite possible that local innovation spillovers do not exist, or are of rather limited scope, and are overestimated in naive regressions that neglect such omitted variable problems.

A way to address this problem is to use a shifter of listed firms' lab activities that is orthogonal with the unobserved local conditions that might affect private firms' innovation. I do so by using the adoption of BC laws and focusing on listed firms' labs outside their state of incorporation and estimate the following equation:

$$\begin{aligned} \text{Log}(Y_{cst}) = & \alpha_c + \delta_t + \beta \text{ Shock Listed Firms}_{ct} \\ & + X_{ct} + \gamma_{st} + \epsilon_{cst}, \end{aligned} \quad (1)$$

where Y_{cst} is the innovative output of private firms located in commuting zone c , state s , and year t . *Shock Listed Firms_{ct}* is defined as: $\sum_i w_{i,0} \times BC_{it}$, namely the weighted average of the adoption of BC law for all firms with a lab in commuting zone c , where the weights $w_{i,0}$ are given

² The data are available at <http://dvn.iq.harvard.edu/dvn/dv/patent>.

Table 1

Summary Statistics.

This table provides summary statistics for the main variables. Statistics have been computed at the commuting zone-Year level. Variables are described in Section 2.

	Mean	Std. Dev.	p(25)	p(50)	p(75)
Patents Private Firms	45	175	1	5	19
Patents Listed Firms	64	269	0	3	17
Stock Listed Patents	293	1,212	2.5	12	78
Population Density	0.41	0.62	0.1	0.21	0.42
Firm Density	0.93	1.5	0.24	0.45	0.92
Share Urban	0.52	0.21	0.37	0.52	0.68
Share Black	0.09	0.12	0.01	0.04	0.12
Share Women	0.51	0.011	0.51	0.51	0.52
Share College Educated	0.39	0.094	0.32	0.39	0.46
Share S&E	0.017	0.0091	0.01	0.01	0.02
Fraction Citation Listed Firms	0.3	0.086	0.26	0.3	0.35
Fraction Citation Local Firms	0.033	0.035	0.01	0.02	0.05
Mobile Inventors from Listed Firms	5	16	0	0	2
Share Inventors Previously in Listed Firms	0.066	0.11	0	0	0.11
Spin outs	3.9	15	0	0	2
# Non Local VC Investments	6.3	49	0	0	0

by the fraction of patents by firm i in the total amount of listed patents at the beginning of the period.³ α_c and δ_t denote commuting zone and year fixed effects respectively. Commuting zone fixed effects capture time-invariant determinants of innovation in each area, such as the geographic characteristics or the presence of an important university. Year fixed effects control for aggregate shocks and common trends in innovation activity produced by legal and institutional changes at the federal level, such as the creation of the Court of Appeals for the Federal Circuit in 1982. Finally, I add state×year fixed effects denoted by γ_{st} to remove any time-varying shocks or state characteristics that might affect innovation by all firms, such as state business cycles, or time-varying state institutional and policy differences (e.g., marginal tax rate).

The parameter of interest is β , which measures the extent to which private firms react to the overall shocks affecting listed firm labs. Given the state×year fixed effects, β only captures spillovers that occur within a state across commuting zones and does not include variation coming from commuting zones in different states. I cluster standard errors at the commuting zone level.

Such a strategy makes it possible to causally estimate the existence of spillovers from listed firm activities on private firms' innovation. It might not pin down the exact elasticity between listed patents and private patents if the adoption of BC law affects other characteristics of listed firms' labs. In other words, it might not be possible to treat *Shock Listed Firms* as an instrument as the exclusion restriction not only requires that the shock is uncorrelated with local productivity shocks (which it is), but also that it affects private firms innovation only through the change in innovation by listed firms and not other listed firm policies. In the case of BC laws, they often lead

to slightly higher wages and a small decline in firm efficiency (Bertrand and Mullainathan, 2003). Those changes may create spillovers onto private firms and affect their innovation policies. While this does not affect the validity of the shock to study how an exogenous shock on listed firms spills over onto local private firms, it does change the interpretation of why private firms' innovation change.

In Section 5, I discuss the extents and limits under which it is possible to consider *Shock Listed Firms_{ct}* as an instrument for listed firms' innovation and to interpret β as the elasticity between private and listed patents.

3.2. Exogenous variation in innovation by listed firms

3.2.1. Antitakeover laws

In the 1980s and early 1990s, US states adopted what are generally referred to as the “second generation” of antitakeover laws. The most stringent of these are called “business combination laws” (BC laws).⁴

BC laws strongly limit the likelihood that a firm will be the target of a highly leveraged hostile takeover, by restricting a raider's ability to sell the assets of the acquired firm. Because these takeovers are frequently financed by means of the sale of certain of the target's assets, BC laws have effectively insulated managers from hostile takeovers by giving management the right to “veto” a takeover by making it more difficult to finance. Therefore, their adoption can be considered as a valid source of variation in corporate governance. In particular, BC laws allow managers to follow preferences that are not necessarily aligned with shareholders' best interests. Two types of these preferences would lead to a decline in innovation. First, managers might exert less effort based on their intention to “enjoy the quiet life” (Bertrand and Mullainathan, 2003). Second, risk-averse or career-concerned managers might undertake less risk than desired by a diversified shareholder

³ As it is common in the literature, I define patents by listed firms as the stock of patents that listed firms have produced in commuting zone c at time t using the standard perpetual inventory method. The stock of listed patents in year t is $Stock_t = (1 - \eta)Stock_{t-1} + Listed\ Patents_t$ where $Listed\ Patents_t$ is the number of new patents filed by listed firms in year t and $\eta=0.15$.

⁴ For a detailed history of first and second generation of antitakeover laws, see Bebchuk et al. (2002).

and decide to “play it safe” (Gormley and Matsa, 2016). Both types of behavior have been found to increase after BC laws were adopted.

One might be concerned that the adoption of BC laws would have a direct impact on both listed and private firms. This is unlikely for two reasons. First, private firms are closely held and therefore have much lower agency problems between the management team and the owner of the firm, because in the vast majority of cases there is no separation between the equity owners and the firm's top management team. As such, managers cannot become more or less insulated from equity holders and agency costs do not vary with the adoption of BC laws. Second, because private firms are closely held, a raider will always have to deal directly with the owner of the firm who can refuse or accept the transaction. The opinions of the non-owner managers are irrelevant and the adoption of BC laws does not change that.

3.2.2. Exogenous shock on innovation by listed firms using BC law adoptions

I start by showing that the adoption of BC laws is an exogenous shifter of innovation by listed firms in the different commuting zones that have labs. After I drop all patents in commuting zones located in the state of incorporation of the firm, I estimate the following equation:

$$\log(1 + \text{ListedPatents}_{it}) = \alpha_i \times \gamma_c + \delta_t + \beta BC_{it} + \epsilon_{it}, \quad (2)$$

where BC_{it} is a dummy variable equals to one if firm i is incorporated in a state that has passed a BC law after year t . $\alpha_i \times \gamma_c$ denotes firm \times commuting zone fixed effects, and δ_t denotes year fixed effects.

Identifying assumptions. Using the adoption of business combination law as a shock to innovation produced by listed firms face two problems. First, the adoption of the law may change or reflect the state's economic context. To deal with this, I exploit the geographic dispersion of innovation by listed firms. For instance, listed firms file only 20% of their patents in their state of incorporation. So I exclude from my analysis innovation by firms in their state of incorporation. For example, I consider a firm incorporated in Virginia but that files patents in Austin, Texas. When Virginia passed a BC law in 1988, the firm reduced its innovation in all areas, including Austin. I use this to study the impact on innovations by local private firms in Austin.

Second, even if the adoption of a BC law constitutes a plausible source of exogenous variation in the number of patents produced by listed firms already located in a given area, one source of endogeneity remains. Indeed, the allocation of where a listed firm decides to conduct its research activity initially is not a random decision. For instance, assume that Austin-San Marcos (Texas) experiences a positive productivity shock that increases the innovation of both listed and private firms. In that event, listed firms are more likely to conduct their research there. The number of listed firms affected by the shock will therefore increase, leading to an increase in the estimated amount of innovation produced by listed firms. The higher volume of innovation by listed firms will be matched by more innovation by private firms, as both types of firms benefit from a positive productivity shock, leading to a spurious correlation

between patents filed by listed firms and patents filed by private firms. However, after the first year, the evolution of patents by listed firms will again only depend on the BC laws. Therefore, the threat to identification comes from the entry (and exit) of listed firms in and out of the sample.

This problem would also arise in context where the econometrician predicts within firm variation with an instrument, but then aggregate the prediction at the industry level, such as in Bloom et al. (2013). To see this, note that I can decompose the growth of total patents over h years g_t^h into five margins:

$$\begin{aligned} g_t^h = & \underbrace{\log\left(\frac{\bar{Pat}_t(I_{t \cap t-h})}{\bar{Pat}_{t-h}(I_{t \cap t-h})}\right)}_{\substack{i_t^h = \text{intensive margin} \\ \text{entry margin}}} \\ & + \underbrace{\log\left(\frac{N_t(I_t)}{N_t(I_{t \cap t-h})}\right)}_{\text{entry extensive}} + \underbrace{\log\left(\frac{\bar{Pat}_t(I_t)}{\bar{Pat}_t(I_{t \cap t-h})}\right)}_{\text{entry intensive}} \\ & - \underbrace{\log\left(\frac{N_{t-h}(I_{t-h})}{N_{t-h}(I_{t \cap t-h})}\right)}_{\text{exit extensive}} - \underbrace{\log\left(\frac{\bar{Pat}_{t-h}(I_{t-h})}{\bar{Pat}_{t-h}(I_{t \cap t-h})}\right)}_{\text{exit intensive}} \end{aligned}$$

where I_t is the set of firms that innovates in year t , $N_t(I_t)$ is the number of firms that innovate in year t , and $\bar{Pat}_t(I_t)$ is the average number of patents per firm during year t . \cap denotes the set of firms that innovate in both year t and in year $t-h$. The total entry and exit margin can be decomposed into the number of firms that enters and exits and the average number of patents filed by entering and exiting firms.

This decomposition highlights that while the firm level instrument will address endogeneities raising from the intensive margin, it does *not* remove the biases coming from the decisions of firms to enter and exit the sample, which are likely correlated with unobserved local productivity shocks.

To address this problem, I focus on listed firms present for the entire sample period and consider that they are present from the beginning in all the commuting zones in which they will patent at some point in time. This ensures that the only variation in patents by listed firms comes from the adoption of the BC law.⁵

Estimation. Table 2 shows the effect of adopting a BC law on listed firms' innovation for the balanced sample from 1975 to 2000. Adopting a BC law generates a decline in patenting between 4 and 6%, depending on the specification, and is always highly significant at the 1% level. Columns 2–5 report the results of various robustness tests. I add industry \times year fixed effects to absorb time-varying fluctuations at the industry level (such as technology or

⁵ One potential problem with this strategy is that it reduces the number of listed firms to 1491 firms. I also run a similar regression with the complete sample (16,914 firms) to generate a prediction based on this sample and find similar results, which suggests that the magnitude of the bias that entry could produce is very small.

Table 2

Effect of BC Laws on Patenting by Publicly Listed Firms.

Dependent variable is the log of patents filed by Compustat firms in a given year and commuting zone for the sample of firms present from 1975 to 2000. The regression for column 2 includes industry×year FE. I then add commuting zone×year FE in column 3. I exclude all firms incorporated in Delaware and all innovation activity in California in columns 4 and 5 respectively. In column 6, the sample is at the firm-year level and R&D is scaled by firm capital. Standard errors are clustered by commuting zone and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance, respectively.

Sample	Patents					R&D
	All			Exc. Delaware	Exc. California	All
	(1)	(2)	(3)	(4)	(5)	(6)
<i>BC Adoption</i>	-0.04*** (0.01)	-0.06*** (0.01)	-0.05*** (0.01)	-0.06*** (0.02)	-0.06*** (0.01)	-0.02** (0.01)
Fixed Effects						
Commuting Zone	✓	✓	—	—	—	—
Firm	✓	✓	✓	✓	✓	✓
Year	✓	—	—	—	—	—
Industry× Year	—	✓	✓	✓	✓	✓
Commuting Zone× Year	—	—	✓	✓	✓	✓
Observations	183,168	183,168	183,168	87,630	169,525	34,550

sale shocks) in column 2. In column 3, I add commuting zone×year fixed effects to absorb any commuting zone-specific time-varying shocks shared by all firms in the same commuting zone, such as localized business cycles or productivity shocks.⁶ I then excludes listed firms incorporated in Delaware and exclude patents filed in California in columns 4 and 5, respectively.

I also check that the results are not capturing a trend by plotting in Fig. 1 the evolution of patenting activity around the regulation date. I estimate Eq. (2) but replace the adoption of the BC law with dummy variables for each year from 10 years before to 10 years after the regulation. Reassuringly, the figure shows that there is no trend before the event date. It also shows that the effect of the regulation materializes only progressively after the event date, which is expected as firms need time to adjust to new environments.

Identifying the effects of innovation by listed firms on private firms exploits the fact that commuting zones will be more or less affected by the shock generated by the adoption of BC laws. Fig. 2 shows the distribution of patents filed by listed firms before 1984 (the last year before the adoption of the first BC law). This figure shows the distribution of patenting activity by listed firms that will be affected at some point in time by BC laws. The figure also shows that listed firms affected by BC laws represent an important part of all patents filed by listed firms throughout the US, which reduces the risks that my estimation will only capture evolution that is specific to a limited number of geographic areas.

Discussion. What margin of adjustment can explain why listed firms experience a reduction in their patent production? A challenge when estimating the effect of the adoption of BC laws on R&D spending is that most R&D spending is actually the wages of employees in-

volved in the research activities of the firm (around 80%–90% of firm total R&D spending). Since Bertrand and Mullainathan (2003) show that managers insulated from takeover risks increase wages for the firm's white collar workers, this might translate into higher R&D “spending,” even though efforts to innovate have actually gone down. Patents are therefore a much better proxy for the knowledge produced by listed firms in this context, as they provide a measure of the amount of knowledge produced by listed firms that can be observed and reused by local private firms. With this caveat in mind, replacing patents by the amount of R&D scaled by firm assets in Eq. (2) yields a negative coefficient of 2 percentage points, significant at the 5% level, which represents a 18% decline relative to the sample mean (column 6 of Table 2). Taken at face value, this implies an elasticity of R&D spending to patent product of around 0.3% (0.06/0.18), meaning that when listed firms reduce their R&D spending by 1% (the input) due to the shock, they reduce the amount of patents produced by 0.3% (the output).

4. Local innovation spillovers

4.1. Baseline estimation

I begin by investigating the effect of a change in innovation by listed firms on the number of patents filed by private firms in a given commuting zone in the following year. The results are reported in Table 3. Column 1 shows the naive OLS from Eq. (1). The elasticity of patents filed by private firms to patents produced by listed firms is 0.24. Columns 2–5 report the effect when I instead use the average shock on listed firms coming from the adoption of BC laws in different states. In every case, the effect is negative and strongly significant at the 1% level, implying that an increase in the number of listed firms' research labs that are shocked, reduces innovation by private firms in the same commuting zone.

I add various controls at the commuting zone level that might affect the propensity of private firms to innovate in columns 3 and 4. In column 3, I add economic and

⁶ For example, assume that I have only two firms in a given commuting zone. The identification comes from the fact that one firm will be incorporated in New York where a B.C. law was adopted in 1985, whereas the other is incorporated in California where such a law has never been adopted.

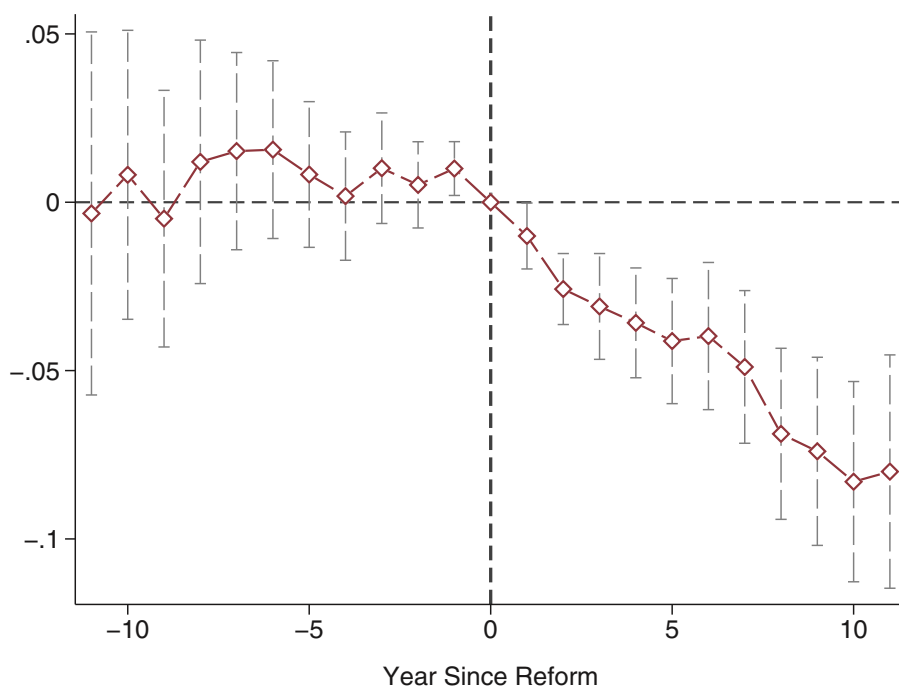


Fig. 1. Effect of BC laws on patenting by publicly listed firms. The figure shows the evolution of innovation around regulation dates. The specification is the same as Eq. (2) except that the dummy for the adoption of business combination law is replaced by a collection of variables $I(k)$, where $I(k)$ is a dummy equal to one exactly k years after (or before if k is negative) the state implements the regulation. The solid line plots the point estimates for $k = -10, \dots, 10$, using the regulation year $k = 0$ as the reference year. The dashed lines plot the 95% confidence interval.

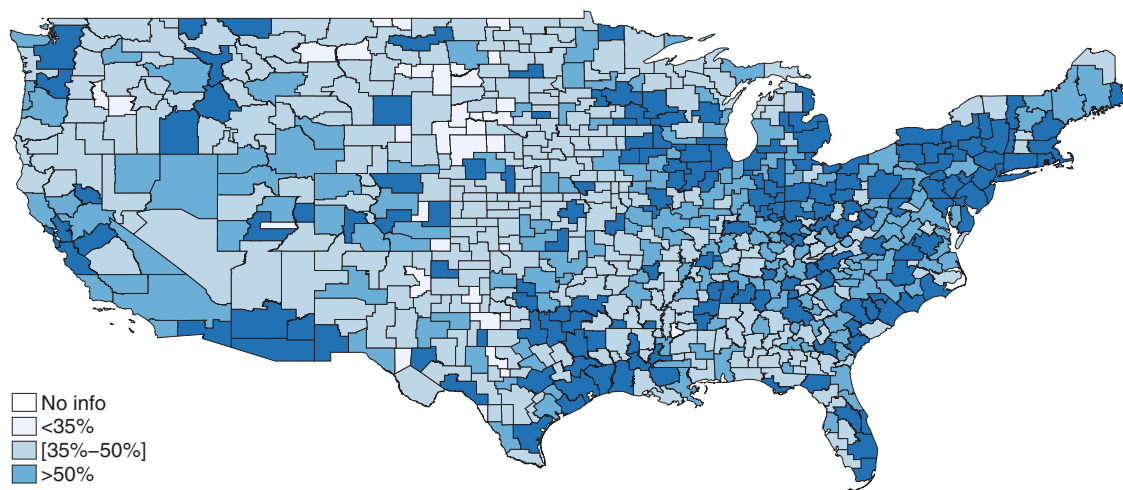


Fig. 2. Fraction of Patents by listed firms Affected by BC Laws. This map shows the geographic dispersion of publicly listed firms that will be affected by the adoption of BC laws. I calculate the fraction of patents filed by listed firms affected over the total of patents filed by listed firms for each Commuting Zone.

demographic controls: share of African-Americans, share of women, population density, share of population living in an urban area, share of self-employed, and industrial specialization.⁷ In the regression for column 4, I add innovation-specific controls: number of doctorates

granted each year, number of existing college institutions reported by the Integrated Postsecondary Education Data System (IPEDS), *Technology specialisation* defined as the local Herfindahl of technology classes (thus in both cases, the greater this measure, the more highly specialized that a given commuting zone is); *Technology age*, defined as

⁷ *Industry specialization* is defined as the local Hirschmann-Herfindahl Index for the 10 economic sectors available in the BEA. Those sectors include the following: agriculture, mining, construction, manufacturing,

transportation, wholesale trade, retail trade, finance, services, public administration.

Table 3

Listed Firm Spillovers on Innovation by Private Firms.

The dependent variable is the log of patents filed by private firms. *Shock* is defined in Eq. (1) as the weighted sum of the adoption of BC laws for listed firms. *Shock Listed Firms-Close CZs* is the average shock affecting listed firms in the four closest commuting zones around private firms and *Shock Listed Firms-Distant CZs* is the average shock affecting listed firms in the next four closest commuting zones. In the regressions for columns 3 and 4, I add various controls at the commuting zone-year level. In column 6, *Predicted Listed Patents* is the log of listed patents predicted by the adoption of BC laws. The regression is estimated with a 2SLS procedure to correct the standard errors. Standard errors are in parentheses and clustered at the commuting zone level and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance respectively.

	(1)	(2)	(3)	(4)	(5)	(6)
Listed Patents	0.201*** (0.021)					
Predicted Listed Patents						0.203*** (0.042)
Shock Listed Firms		-0.097*** (0.020)	-0.092*** (0.020)	-0.095*** (0.020)	-0.093*** (0.020)	
Shock Listed Firms-Close CZs					-0.058* (0.032)	
Shock Listed Firms-Distant CZs					0.033 (0.062)	
Fixed Effects						
Commuting Zone	✓	✓	✓	✓	✓	✓
State × Year	✓	✓	✓	✓	✓	✓
Economic Controls	—	—	✓	✓	—	—
Innovation Controls	—	—	—	✓	—	—
Observations	17,125	17,125	17,125	17,125	17,125	17,125

the average age of technologies exploited in a commuting zone capturing the fact that commuting zones working in newer, more fertile technologies may generate more patents (Hombert and Matray, 2017) and the amount of venture capital investment made. Because several of these variables are likely to be directly influenced by innovation produced by listed firms, I only use demographic controls and the number of establishments in the rest of the paper.⁸

While the point estimate of the coefficient of interest remains relatively stable, it does slightly fluctuate, which might reflect that multiple channels connect the BC laws to private firms' innovative activities, besides the effect of innovation by listed firms. However, coefficient movements alone are not fully informative of the degree of robustness. To quantify the effect to which further unobserved characteristics might drive the results, I compute the bounds in Oster (2019) and obtain a value for the δ parameter of 5.7, well above the recommended value of 1. This implies that it is unlikely that further unobserved characteristics could be driving down the results.⁹

4.2. Effect of distance

In column 5, I explore how the effect evolves with distance. I define *Shock Listed Firms-Close CZ_{ct}* as the sum of patents produced by listed firms in the four closest neighbor surrounding the commuting zone *c*. I also calculate

the sum of listed patents produced in the next four closest neighbors labeled *Distant CZ_{ct}*. I identify close neighbors and distant neighbors by calculating the geographical distance between each commuting zone using the latitude and longitude of each commuting zone centroid. I find in column 5 that the innovation made by listed firms in close neighboring commuting zones has a small positive effect on the innovation by private firms, but the effect becomes indistinguishable from zero for distant neighbors. This sharp decrease with distance is consistent with other papers documenting that “knowledge does not travel well.” It also implies that analyses attempting to estimate spillovers in innovation at the state level are likely to underestimate their existence because they occur on a much smaller scale.

4.3. Discussion of magnitudes

One issue with this reduced form shock is that it makes the economic interpretation of the magnitude complicated. I address this problem by predicting the amount of listed patents using the overall shock faced by listed firms in the commuting zone and use a 2SLS estimation to adjust standard errors. This offers an automatic rescaling of the coefficient β in Eq. (1) and allows an interpretation in terms of elasticity, since I am regressing a log on a log.¹⁰ The point estimate is equal to 0.2, implying that changing the amount of the innovation made by listed firms by 1% changes the number of patents filed by private firms by 0.2%. To have an estimation in terms of patents, I have to

⁸ See Angrist and Pischke (2008) for a discussion about the problems created by “bad controls.”

⁹ In practise, bounds can be approximated using the following equation: $\beta^* \approx \hat{\beta} - \delta \times [\hat{\sigma}_{\beta} - \hat{\beta}] \times \left[\frac{(R^{\max} - \hat{R})}{\hat{R} - \hat{\sigma}_R} \right]$ where “ $\hat{\sigma}$ ” denotes the variable estimated with all the controls and “ $\hat{\sigma}$ ” denotes variables estimated only when *Shock Listed Firms* is included as a control. I set R^{\max} to one, which is the most conservative. δ is estimated assuming that β^* is equal to zero.

¹⁰ It is important to note that while this 2SLS offers a convenient rescaling of the coefficient of interest, the estimation is not an IV per se as the exclusion restriction might be violated. Indeed, the adoption of BC laws might have other effects on listed firms behaviors that directly affect the innovation of private firms. I discuss these possibilities in the next section.

multiply the elasticity by the ratio of the stock of patents filed by private firms over the stock of patents filed by listed firms. It implies that a variation in one patent filed by listed firms generates a similar variation in 0.14 patent filed by private firms. Another possibility is to perform the following thought experiment. The average listed firm's research lab in a commuting zone has a local stock of around 100 patents. If I relocate this activity to a new commuting zone, it will generate around 14 additional patents by private firms, which would move the commuting zone at the 25th percentile of the distribution to the 75th percentile in term of innovation by private firms. This substantial effect could explain why cities and states compete to attract R&D activities.

5. Spillovers of innovation

While the effect of the adoption of these laws on listed firms innovation is particularly striking, the adoption of BC laws can change the “overall efficiency” of listed firms, which, in return, can spill over onto local private firms located in the same area as affected listed firms' research labs. I explore this possibility in [Section 5.1](#). Finding limited support for this possibility, I then unpack these “innovation spillovers” and test two channels: competition in the market for ideas ([Section 5.2](#)) and knowledge diffusion ([Section 5.3](#)).

5.1. Overall decline in dynamism?

The adoption of BC laws might spill over onto private firms via a change in overall decline in listed firm dynamism through three specific channels. First, listed firms become less productive overall ([Bertrand and Mullainathan, 2003](#)), reducing the pressure that private firms face in general. Second, listed firms can be consumers of local private firms. Third, the adoption of BC laws may affect innovation by private firms via the merger and acquisition (M&A) market.

To test if the change in innovation by private firms is just a “side-product” of an overall change in the dynamism of private firms, I look at the number of small firms, employment, and average wages of firms in non-innovative industries. To do so, I use the County Business Pattern data that report employment and total wages at the county-industry level and focus on firms with less than ten employees. I classify an industry as not innovative if it is below the median of R&D spending distribution in Compustat at the three-digit SIC level. Columns 1–3 of Table A.2 in the Appendix show that there is no discernible effect for this group.

The innovation of private firms could still be indirectly affected by a decline in listed firms' overall efficiency if local private firms supply innovation for nearby listed firms. This possibility seems inconsistent with the very fast decline of spillovers with distance shown in column 3 of [Table 3](#), unless I am willing to assume that customers-suppliers are always in the same commuting zone, but never farther apart than around 200 miles. This seems highly at odds with US data. Using the US Census Bureau's

Commodity Flow Survey, [Holmes and Stevens \(2012\)](#) reports that the average distance between suppliers and customers is 529.6 miles and that less than a third of suppliers and customers are closer than 100 miles.¹¹ This “supplier channel” also makes no prediction regarding the role played by channels fostering knowledge diffusion and in particular the mobility of highly skilled workers, something I find strong evidence for. In particular, [Section 5.3.2](#) shows that the spillovers are entirely muted in states where labor regulation highly impedes inventors from moving to different firms in the same state, which should not matter if the spillovers are driven by suppliers-customers link.

In the regression for column 4 in Table A.2, I restrict my estimate to private firms in downstream industries, which are selling products close to final consumers, and remove from the sample all private firms in upstream industries that are more dependent on other firms' demand. I measure the degree of upstreamness using the I-O table for the US and the methodology in [Antràs et al. \(2012\)](#) and consider an industry is upstream if it is above the sample median. I remove from the sample all listed firms that report private firms among their main suppliers in the Compustat Segment data in column 5. In both cases, the point estimate of the variable *Shock Listed Firms* remains virtually unchanged.

The third possible explanation is that the adoption of BC laws affects innovation by private firms via the M&A market. One possibility is that entrepreneurs innovate in order to sell their startup to a large corporation. If the adoption of BC laws reduces listed firms' takeover demand, it might reduce potential targets' incentives to innovate. However, it is unclear why in this case the effect of innovation spillover would be so local or why it would be affected by the degree of inventor mobility. In addition, I estimate whether innovation by listed firms in a given commuting zone affects the likelihood to observe the acquisition of a private firm (column 6) or a private high-tech firm (column 7) in the same commuting zone. I identify the localization of an acquired private firm using SDC Platinum. Similarly, I consider a private firm as “high-tech” if SDC indicates that the firm operates in a high-tech industry. In both cases, I find no effect.

Overall, these results provide evidence that the change in innovation by private firms is driven by a change in innovation by listed firms, rather than a change in other behaviors from listed firms affected by the adoption of BC laws. Since it appears unlikely that the adoption of BC laws affects private firms' innovation via another channel than the innovation of listed firms, in the rest of the paper, I rescale *Shock Listed Firms* by the log of listed firms patents, by creating the variable *Predicted Listed Patents*, such that the effect of the BC law adoption has a direct interpretation as the elasticity between listed firms' innovation and private firms' innovation. Because using directly the pre-

¹¹ The SCF is a survey of the shipments originating in manufacturing, wholesale, and mining establishments. It allows for each firm in the survey to know the exact origin at the ZIP code level of shipments made to the surveyed firm and therefore to construct average distance between customers and suppliers.

dicted variable would underestimate the standard errors, I estimate 2SLS regressions to correct for this problem.

5.2. Competition in the market for ideas

I now examine the channels explaining the existence of these spillovers of innovation. There are two reasons for why a decline in innovation produced by listed firms would reduce innovation by private firms nearby. First, a change in innovation by listed firms can change the degree of competition faced by private firms in the market for ideas, thereby affecting their incentives to innovate. Second, innovation produced by listed firms can expand the amount of knowledge that private firms can use to innovate themselves, in particular since innovation activity generates “knowledge spillovers.”

The effect of competition by innovative listed firms on private firms is a priori ambiguous and will depend on the effect of competition on pre and post-innovation rents. If competition reduces pre-innovation rents, an increase in competition increases firms incentives to innovate to escape the competition (sometimes refers to as the “Arrow effect” (Arrow, 1972)). By contrast, when competition reduces post-innovation rents, higher competition will reduce innovation (sometimes refers to as the “Schumpeter effect” (Arrow, 1972)). Hombert and Matray (2018) show that when innovation improves firm product differentiation rather than firm productivity, the Arrow effect should dominate. They then estimate empirically that the return to innovate increases with product market competition. If I assume that lower competition reduces incentives to innovate, the competition channel would be consistent with the existence of spillovers of innovation. However, to explain these spillovers, it will have to be the case that not only listed and private firms compete in the same markets, but also that the direct competitors of private firms are in the same commuting zone, since I find that innovation spillovers decline quickly with distance. There are reasons to doubt that listed and private firms are often competing neck to neck. For instance, Holmes and Stevens (2014) using US census data show that large and small firms in the same narrowly defined industry are unlikely to compete with each other and that instead are performing different functions.

To test directly if the colocation of competitors can explain innovation spillovers, I run four tests in Table 4. First, I restrict the sample to firms in the tradable sector, as these firms are more likely to compete with firms across the country. Second, I restrict to industries that are more open to international trade, as a proxy for the degree to which firms compete not only across the country but also across the world. To do so, I estimate the ratio of imports plus exports over value of domestic shipments and restrict the analysis to industries with a value above the median.¹² Third, I compute the geographical concentration of patents for all technological classes (industries) and restrict the

analysis to industries that are not concentrated geographically. Concretely, I compute a geographical Herfindahl index for each technology class based on the share of a technology's patent that fall in each commuting zone. I then remove from the sample technologies with a concentration above the sample median. This implies that each innovator in the sample now has competitors in multiple commuting zones. Fourth, I remove private firms that are innovating in the same technological classes as the listed firms present in their commuting zone.

In all cases, the point estimate remain stable and very close to the one estimated using the whole sample. In particular, the regression in column 4 forces the estimation to be made by looking at private firms that are not competing with listed firms. While the results may at first appears puzzling, it is important to remember that a large fraction of citations occur across very broad technological classes.¹³ The fact that private firms innovating in computer industry benefit from the innovation of listed firms that innovate in the pharmaceutical or chemical industries, suggests that the diffusion of knowledge plays an important role in explaining innovation spillovers.

5.3. Knowledge diffusion channels

How and why does knowledge spread locally? In this section, I explore two channels through which knowledge diffuses from innovative listed firms to other private firms in the same area: learning across local firms and inventors moving across existing firms or founding or joining local spin-outs.

5.3.1. Effect depending on learning opportunities

Technological proximity. To test whether the magnitude of innovation spillovers varies with the degree technological proximity between listed and private firms in the same area, I build two proxies for the potential of learning and re-estimate Eq. 1 by interacting the patents produced by listed firms with these proxies. I demean all the proxies and interact them with the main variable *Listed Patents* to obtain the marginal additional effect that each proxy create with respect to the mean effect of *Listed Patents*.

First, I use the propensity of private firms to build on the knowledge produced by local listed firms with patent citations. I measure the fraction of listed patents cited by private firms in a given area over the total of all the citations made by all firms.

Second, I measure technology overlap following Jaffe (1986). For each commuting zone, I calculate the number of patents granted to each firm by technological categories, using the disaggregated three-digit (425 distinct) technological categories. The share of patents granted to firm i located in commuting zone c in each technological class s ($s=1, \dots, 425$) is then arranged in a vector $T_{ic} = (T_{ic1}, \dots, T_{ic425})$. The technological proximity in commuting zone c is defined as the uncentered correlation

¹² imports and exports at the sic-year level come from Pete Schott website and value of domestic shipment from the NBER manufacturing dataset. Because custom data only record flows of good, this test is restricted to manufacturing firms.

¹³ Jaffe et al. (1993) report that up to 25% of citations occur across five broad technological fields. When looking at the three-digit level (approximately 450 technological fields) approximately 40% of citations are across fields.

Table 4

Innovation Spillovers: Competition with Listed Firms.

The dependent variable is the log of patents filed by private firms. In all regressions, *Predicted Listed Patents* is the log of listed patents predicted by the adoption of BC laws. The regression is estimated with a 2SLS procedure to correct the standard errors. In column 1, I restrict to firms in tradable industries. In column 2, I restrict to manufacturing industries with a value of openness ($[\text{import} + \text{export}] / \text{domestic production}$) above the median sample. In column 3, I drop industries with a geographical HHI of their patents above the sample median. In column 4, I drop private firms that patent in the same technologies than listed firms in the commuting zone. Standard errors are clustered by commuting zone and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance respectively.

Sample	Tradable (1)	High international openness (2)	Low geographic concentration (3)	Non-overlapping technologies (4)
<i>Predicted Listed Patents</i>	0.21*** (0.04)	0.19*** (0.04)	0.23*** (0.04)	0.20*** (0.04)
Fixed Effects				
Commuting Zone	✓	✓	✓	✓
State × Year	✓	✓	✓	✓
Commuting Zone Controls	✓	✓	✓	✓
Observations	17,125	17,125	17,125	17,125

Table 5

Innovation Spillovers Depending on Technology Proximity.

The dependent variable is the log of patents filed by private firms. In all regressions, *Predicted Listed Patents* is the log of listed patents predicted by the adoption of BC laws. The regression is estimated with a 2SLS procedure to correct the standard errors. Each column interacts *Predicted Listed Patents* with a proxy of technology proximity. Column (1) uses the fraction of citations of patents by listed firms made by private firms. Column (2) uses the degree of overlap in technological classes based on the procedure developed by Jaffe (1986). Column (3) uses the degree of proximity across technological classes based on the Mahalanobis distance defined by Bloom et al. (2013). Column (4) uses proxies in columns (1) and (2). Column (5) uses proxies in columns (1) and (3). Standard errors are clustered by commuting zone and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance respectively.

	(1)	(2)	(3)	(4)	(5)
<i>Predicted Listed Patents</i>	0.17*** (0.04)	0.17*** (0.04)	0.19*** (0.04)	0.16*** (0.04)	0.18*** (0.04)
<i>Predicted Listed Patents</i> × Tech. Prox. (Citation Listed Firms)	1.92*** (0.38)			1.61*** (0.38)	1.63*** (0.35)
<i>Predicted Listed Patents</i> × Tech. Prox. (Jaffe Distance)		0.61*** (0.11)		0.50*** (0.11)	
<i>Listed Patents</i> × Tech. Prox. (Mahalanobis Distance)			0.04*** (0.01)		0.03*** (0.01)
Fixed Effects					
Commuting Zone	✓	✓	✓	✓	✓
State × Year	✓	✓	✓	✓	✓
Commuting Zone Controls	✓	✓	✓	✓	✓
Observations	17,125	17,125	17,125	17,125	17,125

coefficient between the vectors of all firm i, j pairings, calculated as: $TECH\ CORR_c = (T_{ic}T'_{jc}) / [(T_{ic}T'_{ic})^{1/2}(T_{jc}T'_{jc})^{1/2}]$. The index ranges from zero to one, depending on the degree of technological overlap between firms. The closer this index is to one, the more that firms located in commuting zone c overlap in technological classes. One drawback of the Jaffe distance is that it considers proximity only within the same technology class. I correct for this problem by using the Mahalanobis distance developed by Bloom et al. (2013), which allows me to calculate a degree of technological proximity between different technology classes.

The correlation between the technological proximity measured by the propensity to cite patents by listed firms and the two other proxies based on technological overlap across patent classes is quite low (between 20% and 30%), implying that these separate proxies capture the different dimensions of technological proximity.

Table 5 reports the results. Column 1 shows the results of the interaction with the propensity of private firms to cite listed firms' patents. Consistent with the intuition that

spillovers should be more important when private firms rely more on knowledge produced by listed firms, I find that the coefficient for the interaction term is positive and strongly significant. In terms of economic magnitude, increasing the fraction of citations of listed firms' patents by one standard deviation increases innovation spillovers by a factor of two. The results in columns 2 and 3 show a similar amplification when I interact listed patents with the degree of technological proximity using the Jaffe distance and the Mahalanobis distance. Finally, the regressions for columns 4 and 5 include two different measures of proximity (citations of listed firms and Jaffe distance or citation and Mahalanobis distance) and the results show that each has a positive impact on spillovers. These results confirm that each measure captures a different dimension of learning opportunities that matters for local innovation spillovers.

Density of skilled workers. Marshall (1890) is among the first to notice that social interactions among workers create learning opportunities that enhance their productivity. As he writes in his *Principles of Economics*: "(...) so

Table 6

Innovation Spillovers Depending on Skilled Worker Supply.

The dependent variable is the log of patents filed by private firms. In all regressions, *Predicted Listed Patents* is the log of listed patents predicted by the adoption of BC laws. The regression is estimated with a 2SLS procedure to correct the standard errors. In column 1, I report the effect when the patents by listed firms is interacted with the supply of scientists and engineers (S&E) in a given commuting zone-year. In column 2, I use the supply of college graduates in a given commuting zone-year. In columns 3 and 4, I instrument the supply of college graduate. The instrument is the share of 15–19 year-old enrolled in school in 1880, constructed from the US Census of 1880 in column 3. In column 4, the instrument is a dummy equal to one if the commuting zone contained a college created via the “Land Grant Movement” in 1862 and 1890 (Nervis, 1962). Standard errors are clustered by commuting zone and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance respectively.

	(1)	(2)	(3)	(4)
<i>Predicted Listed Patents</i>	0.18*** (0.04)	0.18*** (0.04)	0.20*** (0.04)	0.17*** (0.04)
<i>Predicted Listed Patents</i> × <i>S & E Supply</i>	0.05*** (0.01)			
<i>Predicted Listed Patents</i> × <i>College Graduate</i>		0.91*** (0.16)		
<i>Predicted Listed Patents</i> × <i>College Graduate (IV 1)</i>			0.89** (0.38)	
<i>Predicted Listed Patents</i> × <i>College Graduate (IV 2)</i>				1.12*** (0.36)
Fixed Effects				
Commuting Zone	✓	✓	✓	✓
State × Year	✓	✓	✓	✓
Commuting Zone Controls	✓	✓	✓	✓
Observations	17,125	17,125	17,125	17,125
F-test(Enrollement 1880)	—	—	13	—
F-test(Land Grant)	—	—	—	48

great are the advantages which people following the same skilled trade get from near neighborhood to one another. The mysteries of the trade become no mysteries; but are as it were in the air, and children learn many of them unconsciously.”

The challenge with this channel is that economists cannot directly observe communication, discussions or gossip among workers. Instead, I exploit the prediction that spillovers should be more important in areas in which workers can interact and learn more easily from one another. In particular, I expect two commuting zones facing the same shock on listed firms’ innovation to react differently depending on the density of skilled workers.

I construct two measures of skilled workers: the fraction of scientists and engineers and the fraction college graduates in a given commuting zone at the beginning of the sample period. I use the US 1970 census data and aggregate Census Public Micro Samples at the commuting zone level.

Table 6 shows how the density of skilled workers in a commuting zone affect the magnitude of local innovation spillovers generated by listed firms. Consistent with the intuition that having a greater “brain density” fosters local innovation spillovers, I show in column 1 that innovation by listed firms has a greater effect when the supply of scientists and engineers is higher. Because all my proxies are time invariant, the simple term is absorbed by the commuting zone fixed effect. Column 2 shows a similar result when I proxy learning opportunities using the supply of college graduates. The effect is economically sizable and implies that the last quartile of the college graduate distribution experiences spillovers that are twice as large as those experienced by commuting zones in the first quartile of the distribution.

The inherent limit of cross-sectional tests is that, unobserved characteristics may be correlated with the variables used in the cross-section. For instance, commuting zones with a higher supply of college graduates might also differ in other dimensions, such as investment opportunities that could also foster local innovation spillovers. Ideally, I would like to instrument every variable. Although I cannot (unfortunately) find different instruments for each variable, studies on agglomeration economics suggest two possible instruments for the share of college graduates.

The first instrument builds on Beaudry et al. (2010) and uses the share of 15–19-year-olds enrolled in school in 1880, which proxies for the local availability of high schools at that time. To provide a valid instrument, this deep lagged variable must be uncorrelated with current local economy specialization and technology development, which would not be the case if school enrollment in 1880, for instance, was correlated with physical capital at that time and if capital has built up over time. In this case, capital accumulation would make the area more productive, violating the exclusion condition of the instrument. Beaudry et al. (2010) argues that capital and skill were more substitutes than complements prior to the twentieth century. Therefore, the reasons why some areas had better high schools in 1880 were unlikely to be related to economic and technological development in 1880 and in the following periods.

High school enrollment in 1880 is a good predictor of the share of educated workers more than a century later with an F-stat of 13. Column 3 of Table 6 reports the effect of increasing the share of the college-educated population on the magnitude of local innovation spillovers when I instrument *College Graduate* by *School Enrollment 1880*. Again, I find a positive effect, with a similar order of magnitude.

The second instrument uses the presence of college and universities created in the nineteenth century following the “land-grant movement,” which still strongly predicts cross-sectional variation in college share today. Following two acts in 1862 and 1890, the federal government gave every state a grant to establish colleges, which resulted in the creation of 69 colleges and universities, with each state having at least one. Because this program was undertaken well over a century ago and was not dependent on natural resources, land-grant institution is unlikely to be correlated with unobservable factors that affect innovation today.

Using the list of all land-grant institutions provided in the Appendix of [Nervis \(1962\)](#), I create a dummy variable *Land-Grant* which is equal to one if the commuting zone contains at least one land-grant institution. I end up with 63 distinct commuting zones with at least one land-grant institution (in only six cases does the commuting zone contains two land-grant institutions). When I regress the average share of college graduates over the sample period on the *Land-Grant* dummy, I obtain a very significant effect, with a F-stat of 89.

Column 4 shows the result when I instrument *College Graduate* by *Land-Grant* and confirms again that increasing the share of college graduates (in this case because the commuting zone has one land-grant institution) increases the innovation spillovers generated by listed firms.

5.3.2. Local inventor mobility and spin-outs

The second channel through which knowledge can be transferred locally from one firm to another is by inventors moving across firms in the same area. New workers can share ideas regarding how to organize research production, as well as information about new technologies or about failed experiments that they experienced with previous employers.

I use two strategies to test this channel. First, I build on studies examining the effects of “Non Compete Covenants Law.” These laws restrict intrastate job mobility, because they specify a period during which employees cannot take a job with a competing company (typically within the same industry) located in the same state. By affecting the mobility rate of employees, non-compete laws should affect the speed at which knowledge diffuses locally (e.g., [Stuart and Sorenson, 2003](#); [Jeffers, 2019](#)).

I create two measures of state-level differences in enforcing non-compete covenants. The first follows [Stuart and Sorenson \(2003\)](#) and is a dummy variable *Presence of Non-Compete Laws*: this variable equals one if the state enforces non-compete covenants. The second follows [Garmaise \(2009\)](#) and is an index ranging from 0 to 7 that counts the number of employer-friendly provisions: higher values indicate stronger enforceability of non-compete laws. Therefore, an increase in *Intensity of Non-Compete Law* implies that employees will have greater difficulty moving from one firm to another.

I then interact each variable with the patents produced by listed firms. I expect that if knowledge is diffused by labor mobility, more stringent non-compete laws should limit local innovation spillovers.

[Table 7](#) shows that the magnitude of spillovers is affected by non-compete laws. Column 1 reports the re-

sults when I use the dummy variable *Presence of Non-Compete Laws*. Being in a state that enforces non-compete covenants reduces innovation spillovers by 0.8, which is nearly half of the average effect. In columns 2 and 4, I exclude California from the sample because cities in California are characterized by a higher rate of mobility of high-skilled workers than cities in other states and are also more innovative. I find a slightly stronger effect. Column 3 shows the results when I use the degree of enforceability of non-compete laws and confirms that enforcement of non-compete covenants (an increase in the index) limits knowledge diffusion locally by reducing mobility, which ultimately reduces local innovation spillovers. The point estimate of the interaction term is equal to -0.04, which implies that an increase in the enforcement of non-compete covenants strongly reduces local innovation spillovers. Taken together, these results suggest that states can have an important impact on the ability for local agglomerations to generate innovation spillovers by affecting the rate of mobility across local firms.

The second strategy to identify whether local innovation spillovers are the result of inventors moving across firms in the same area is to estimate directly whether variation in innovation by listed firms affects the number of mobile inventors within a commuting zone. To perform this estimation, I use the unique inventor identifier provided by [Lai et al. \(2009\)](#) that permits me to track inventors across firms and ZIP codes.

To measure inventor movement across local firms, I follow papers such as [Marx et al. \(2009\)](#) or [Hombert and Matray \(2017\)](#) and identify an inventor as changing employers when she files two successive patent applications that are assigned to different firms. Because I am interested in innovation spillovers in a given commuting zone from listed firms to private firms, I define an inventor as moving if: (i) the inventor's employer is different from the previous employer, (ii) the current employer is a private firm and the former employer is a listed firm, and (iii) the inventor was working in the same commuting zone.

I construct the following three measures: *# Mobile Inventors from Listed Firms_{ct}* is the number of inventors who work in a private firm at year t in commuting zone c , but who previously worked in a listed firm located in the same commuting zone. *Share of Mobile Inventors from Listed Firms_{ct}* is the fraction of mobile inventors who worked in a listed firm located in the same commuting zone over the total of all mobile inventors who arrive in private firms in year t in commuting zone c ; and *Share Inventors Previously in Listed Firms_{ct}* is the share of all inventors working for private firms in year t in commuting zone c that formerly worked for a listed firm located in the same commuting zone.

I also explore a specific type of inventor mobility: entrepreneurial spin-out. In this case, inventors formerly employed by a listed firm may decide to leave their employer, to join a newly founded local spin-out in which they can exploit their knowledge and experience. I define a spin-out as follows. Using the unique firm identifier in the NBER patent data, I identify first all the new private firms that appear in the database. Then, I look at all the inventors who work for a new firm in the first year after it appears.

Table 7

Innovation Spillovers Depending on Non-Compete Laws.

The dependent variable is the log of patents filed by private firms. In all regressions, *Predicted Listed Patents* is the log of listed patents predicted by the adoption of BC laws. The regression is estimated with a 2SLS procedure to correct the standard errors. All regressions include commuting zone, year and state×year fixed effects. In column 1, I report the effect when patents by listed firms are interacted with a dummy indicating whether the commuting zone is in a state that enforce non-compete covenants (Stuart and Sorenson, 2003). In columns 2 and 4, I exclude California. In column 3, I use the degree of enforceability of non-compete laws as an interaction term reported in Garmaise (2009). Standard errors are clustered by commuting zone and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance respectively.

Sample	All (1)	Exc. California (2)	All (3)	Exc. California (4)
<i>Predicted Listed Patents</i>	0.21*** (0.03)	0.21*** (0.03)	0.32*** (0.07)	0.34*** (0.07)
<i>Predicted Listed Patents</i> × <i>Presence of Non-Compete Law</i>	-0.08* (0.04)	-0.09** (0.04)		
<i>Predicted Listed Patents</i> × <i>Intensity of Non-Compete Law</i>			-0.04** (0.01)	-0.04*** (0.01)
Fixed Effects				
Commuting Zone	✓	✓	✓	✓
State× Year	✓	✓	✓	✓
Commuting Zone Controls	✓	✓	✓	✓
Observations	17,125	16,675	17,125	16,675

Table 8

Effect on Inventor Mobility from Listed Firms to Private Firms.

This table shows the mobility of inventors to private firms within the same commuting zone. In all regressions, *Predicted Listed Patents* is the log of listed patents predicted by the adoption of BC laws. The regression is estimated with a 2SLS procedure to correct the standard errors. In column 2, I report the fraction of mobile inventors who come from listed firms over the total of mobile inventors to private firms. In column 3, I use the fraction of inventors currently employed by private firms who formerly worked for a listed firm in the same commuting zone. In column 4, I report the number of spin-outs (defined as new private firms employing, in the first year they file patents, inventors formerly employed by a listed firm in the same commuting zone). Standard errors are clustered by commuting zone and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance respectively.

	# Mobile Inventors from Listed Firms (1)	Share Mobile Inventors from Listed Firms (2)	Share Inventors Previously in Listed Firms (3)	# Spin-outs (4)
<i>Predicted Listed Patents</i>	0.06*** (0.02)	0.05*** (0.01)	0.02*** (0.00)	0.09*** (0.02)
Fixed Effects				
Commuting Zone	✓	✓	✓	✓
State× Year	✓	✓	✓	✓
Commuting Zone Controls	✓	✓	✓	✓
Observations	17,125	17,125	17,125	17,125

If at least one of the inventors formerly worked for a listed firm in the same commuting zone, I consider the new firm to be a spin-out. I end up with 22,627 spin-outs, which represents 20% of the total of new firms I observe in the patent data.

Table 8 shows how innovation by listed firms in a given commuting zone can affect inventor mobility flow from listed firms to private firms in the same area. The results in column 1 show that an increase in listed patents generates a higher number of inventors who move from listed firms to private firms. In column 2, I show that this effect is not simply due to an increase in overall mobility, but that inventors formerly working for listed firms represent a higher fraction of mobile inventors who come to work at a new private firm. Column 3 shows that inventors who formerly worked for a listed firm represent an increasing fraction of inventors employed by private firms. In terms of magnitude, doubling patents by listed firms increases the share of inventors employed by private firms who formerly worked for listed firms by 50%. Finally, column 4 shows that spin-out creation in the commuting zone increases

with patents produced by listed firms locally, which provides direct evidence that local innovation spillovers are produced in part because former employees join spin-outs created in the same area and benefit from the knowledge produced in their new employee's previous firm.

6. Local innovation spillovers and venture capital

In this section, I investigate how local innovation spillovers interact with investment by VC funds. If innovation by listed firms active in a commuting zone fosters innovation by local private firms, VC funds from outside the commuting zone should be expected to invest more in the local area where those innovation spillovers occur. Conversely, capital availability should affect the magnitude of local innovation spillovers as private firms are likely to have credit constraints.

Table 9

Capital Inflow: Investments by Non-Local VC Funds.

In all regressions, *Predicted Listed Patents* is the log of listed patents predicted by the adoption of BC laws. The regression is estimated with a 2SLS procedure to correct the standard errors. In columns 1–2, the dependent variable used in the regressions is the number of VC investments made by non-local VCs. In columns 3–4, I examine the total amount invested by non-local VC funds. In columns 2 and 4, I exclude from the sample commuting zones considered as VC centers. All dependent variables are in log. Standard errors are clustered by commuting zone and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance respectively.

# Investments	Total Value			
	(1)	(2)	(3)	(4)
<i>Predicted Listed Patents</i>	0.045** (0.022)	0.045** (0.021)	0.227** (0.097)	0.218** (0.097)
Fixed				
Effects Commuting Zone	✓	✓	✓	✓
State × Year	✓	✓	✓	✓
Commuting Zone Controls	✓	✓	✓	✓
Observations	17,125	17,125	17,125	17,125
Sample	All	Exc. VC centers	All	Exc. VC centers

6.1. Capital inflows

To study VC funds investments geographically, I use VentureXpert, which records both the geographic localization (ZIP code) of the VC fund and the localization of the company in which the fund makes an investment. This allows me to identify precisely when and where investments are made and whether the investments come from a fund located in a different area.

I use two different proxies for the ability of commuting zones to attract out-of-town VC money: the number of investments made and the total value of all investments made in a given commuting zone-year. Each variable is in logs and calculated only for non-local VC funds.

Because the VC industry is highly clustered in three metropolitan areas (combined statistical areas or CSAs) in the US (San Francisco/San Jose, Boston, and New York). I estimate the different models on the entire sample and then I exclude the 16 commuting zones that belong to these three areas.

Column 1 of Table 9 shows that patents filed by listed firms in a given commuting zone in the previous years increase the likelihood that this commuting zone attracts investment from VC funds located in other commuting zones. Column 3 shows similar results when I use the total money invested in a commuting zone-year. Columns 2 and 4 report that the effects are similar when I exclude “VC centers” from the sample.

This result is notable because non-local investments are rare in the VC industry (Chen et al., 2010). Indeed, VC firms interact frequently with companies in which they invest, they monitor and coach the management team (e.g., Lerner, 1995).

6.2. Effect depending on capital availability

In this section, I investigate whether venture capital availability influences the magnitude of spillovers, something that has received little attention in the literature thus far. To do so, I interact the variable *Listed Patents* with the total amount of investments made by VC funds.

To generate exogenous variation in the local availability of capital, I build on studies showing that public pen-

Table 10

Innovation Spillovers Depending on Fund Availability The dependent variable is the log of patents filed by private firms. In all regressions, *Predicted Listed Patents* is the log of listed patents predicted by the adoption of BC laws. The regression is estimated with a 2SLS procedure to correct the standard errors. In column 1, I report the effect when the patents produced by listed firms is interacted with the amount of VC investments made in the state (in log and demean to restore main effects). In column 2, I instrument the amount of VC investments using the value of assets held by local and state pension funds. In the first stage, the coefficient on this variable is 0.30 with an F-statistic of 30. In column 3, I exclude from the sample commuting zones considered as VC centers. Standard errors are clustered by commuting zone and reported in parenthesis. *, **, and *** denote 10, 5, and 1% statistical significance respectively.

	(1)	(2)	(3)
<i>Predicted Listed Patents</i>	0.16*** (0.02)	0.17*** (0.02)	0.17*** (0.09)
<i>Predicted Listed Patent</i> × VC	0.02*** (0.00)		
<i>Predicted Listed Patent</i> × VC (IV)		0.05*** (0.01)	0.05*** (0.01)
Fixed Effect			
Commuting Zone	✓	✓	✓
State × Year	✓	✓	✓
Commuting Zone Controls	✓	✓	✓
Observations	17,125	17,125	17,125

sion funds display a “home-bias” and are more likely to invest the asset under their management in local private equity funds or venture capital funds. As a result, fluctuations in public pension assets in the home-state of VC funds will affect the ability of domestic VC funds to raise capital, which will generate variation in the amount of money they can invest.

I obtain data for state public pensions from the State and Local Government Public-Employee Retirement Systems annual survey conducted by the US Census Bureau and available since 1970. I compute the amount of asset holdings of the state pension fund for every year and use it as the instrument for the total amount of VC investments made at the state level.

Table 10 reports the results for the different proxies. In the regression for column 1, I use the volume of investments made by VC funds in log in a given state-year and

find that greater levels of VC investment increase the magnitude of local innovation spillovers. The coefficient for the interaction term is positive and statistically significant at the 1% level. However, because VC investments are likely to be endogenous with innovation activity realized by listed firms, I instrument VC investments by the amount of local and state public pension funds in the regression for column 2. The first stage produces an F-stat of 30. The IV estimate yields similar results and shows that exogenous variation in the amount of available VC capital amplifies local innovation spillovers. The magnitude of the amplification is important because moving from the 25th percentile to the 75th percentile increases the elasticity by more than 0.4, which is twice the size of the average effect. I reproduce the analysis when I exclude those commuting zones belonging to a “VC center” and show a similar effect in column 3.

Overall, these results demonstrate that capital typically relocates to areas in which local innovation spillovers occur and that in return, capital availability amplifies the magnitude of these local innovation spillovers. This suggests that capital mobility can contribute to increase the differences between geographic entities rather than to narrow such differences.

7. Robustness

In Table A.3, I explore the robustness of my main result. I add a specific technological trend at the commuting zone level to my main specification. Differences in sectoral growth rates or changing propensities to seek patents might affect my findings if for instance, the commuting zones in which patents by listed firms increase more are simultaneously initially more specialized in a growing sector. I thus include a measure of expected commuting zone-level patenting based on pre-period technological specialization and national patenting trends. To predict patenting growth based on initial specialization, I calculate the initial innovation specialization using the 37 different “technological subcategories” (variable *subcat* in the NBER Patent database) and interact this specialization with the aggregate patenting growth of each in each of the 37 categories. I interact the variable with a time trend and add it as a control. In the regressions for columns 2–4, I exclude various commuting zones/firms. In the regressions for columns 2 and 3, I exclude various commuting zones to ensure that my estimate does not reflect the specificity of certain cities (and in particular the most innovative ones). In column 2, I exclude all the commuting zones that belong to one of the five main high-tech clusters: Austin, Boston, Raleigh-Durham, San Diego, and Silicon Valley (namely San Francisco-Oakland-San Jose). In column 3, I directly exclude all the commuting zones within California and Massachusetts, which are considered the two most innovative states. In both cases, the estimates are similar to the initial result. Finally, in the regression for column 4, I exclude patents by listed firms that are incorporated in Delaware and column 5 exclude patents that are filed in commuting zones located in the state in which the listed firm has its headquarters. Again, my results remain unaffected.

8. Conclusion

Using a novel strategy to generate local shocks on the innovation activities of listed firms, I provide evidence for the existence of complementarities between the innovation of listed firms and private firms. Those complementarities explain why a shock on the innovation production of some firms can transmit to the rest of the local economy, although other firms are not directly hit by the shock.

I then explain these complementarities with local information transmission and identify different channels through which this transmission may occur. In particular, the ease with which workers can exchange ideas and learn from one another, the possibility for workers to move from one firm to another, or to create their own firms are all channels through which knowledge is transmitted within the local area. Those results also suggest that state policies can play an important role in affecting the magnitude of local innovation spillovers by shaping the ability for local markets to absorb new knowledge and affecting labor mobility.

Finally, I find that local innovation spillovers generated by listed firms induce venture capital funds from outside the area to invest more into areas where local innovation spillovers happen because these places typically become more productive. I also find that variation in the amount of capital available amplifies the magnitude of innovation spillovers. This last result suggests that finance could be an important factor for explaining the important disparities between cities in terms of economic specialization, entrepreneurship and growth, etc. If capital follows innovation and in return magnifies economic spillovers, small differences between areas can become rapidly amplified.

Assessing exactly and to what extent capital flow is responsible for how agglomerations are formed, sustained, and strengthened offers interesting avenues for future research.

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